

Berry-Esseen bounds for general nonlinear statistics, with applications to Pearson's and non-central Student's and Hotelling's

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Main results

- Initial motivation and general idea
- Application: a BE bound for Pearson's R
- Applications: uniform and nonuniform BE bounds for Student's t and self-normalized sums

Supporting and/or related results

- Comparison: self-normalized sums vs. Student's statistic
- Exact bounds on the closeness between t_p and $t_\infty = N(0, 1)$
- Tail monotonicity of t_p
- Exact upper bounds on the mean and exact lower bounds on the exponential moments of the Winsorised-tilted distribution
- An asymptotically Gaussian bound on the Rademacher tails
- Refined and generalized Bennett-Hoeffding bound
- Positive-part moments via the Fourier–Laplace transform
- Improved & generalized von Bahr–Esseen inequality and applications to concentration of measure for separately Lipschitz functions on product spaces
- Optimal re-centering inequality

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- ▶ Needed: closeness to normality uniformly near H_0 .
- ▶ Kendall's and Spearman's coeffs. are U -statistics, with known BE bounds.
- ▶ For Pearson's, a BE bound is not found in literature.
Hardly surprising:
 - ▶ an optimal BE bound for Student's t : obtained only in '96, by Bentkus and Götze;
 - ▶ A necessary and sufficient condition, in the i.i.d. case, for t to be asymptotically normal: obtained only in '97, by Giné, Götze and Mason.

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4/23 Pearson's R : general idea

Let $(Y, Z), (Y_1, Z_1), \dots, (Y_n, Z_n)$ be i.i.d. random points in \mathbb{R}^2 ; w.l.o.g. $E Y = E Z = 0$ and $E Y^2 = E Z^2 = 1$.

$$\text{Pearson's } R := \frac{\overline{YZ} - \bar{Y} \bar{Z}}{\sqrt{\overline{Y^2} - \bar{Y}^2} \sqrt{\overline{Z^2} - \bar{Z}^2}} = f(\bar{V}),$$

where $\bar{V} := \frac{1}{n} \sum_{i=1}^n V_i$ and the V_i 's are iid copies of $V := (Y, Z, Y^2 - 1, Z^2 - 1, YZ - \rho)$, with $\rho := E YZ = \text{Corr}(Y, Z)$, so that $E V = 0$.

So, $f(\bar{V}) \approx f(0) + L(\bar{V})$, where $L := f'(0)$ is a linear functional.

From here, using exp. ineqs. for sums in B -spaces by Pinelis–Sakhanenko '85: BE bounds $O(n^{-1/2} \ln^{3/2} n)$ if $\|V\|_3 < \infty$, and $O(n^{1-p/2})$ if $\|V\|_p < \infty$ for some $p \in (2, 3)$.

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5/23 The Chen and Shao method '01–'07: a concentration method in Stein-type framework

Given:

- ▶ an abstract nonlinear statistic T ;
- ▶ an abstract linear statistic W .

For $\Delta := T - W$, start with

$$-\mathbb{P}(z - |\Delta| < W \leq z) \leq \mathbb{P}(T \leq z) - \mathbb{P}(W \leq z) \leq \mathbb{P}(z < W \leq z + |\Delta|).$$

A number of applications were given by Chen and Shao.

We modify their method, apply it to $f(\bar{V}) \approx f(0) + L(\bar{V})$, and use other tools to get:

BE-type uniform and nonuniform bounds for statistics of the form $f(\bar{V})$;

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6/23 A resulting corollary for Pearson's R :

If e.g. $\rho = E YZ = 0$ and $\tilde{\sigma} := \sqrt{E Y^2 Z^2} \neq 0$ then

$$\left| \mathbb{P}\left(\frac{R}{\tilde{\sigma}/\sqrt{n}} \leq z\right) - \Phi(z) \right| \leq \frac{4.08}{\sqrt{n}} (\|Y\|_6^6 + \|Z\|_6^6) (1 + \tilde{\sigma}^{-3}).$$

Recall:

here $V = (Y, Z, Y^2 - 1, Z^2 - 1, YZ)$;

so, $\|V\|_3^3 \asymp \|Y\|_6^6 + \|Z\|_6^6$.

Cf. Bhattacharya–Ghosh '78, Chibisov '80: Asymptotic expansion for the distr. of statistics admitting a stochastic expansion.

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$$\left| \mathbb{P}\left(\frac{R}{\tilde{\sigma}/\sqrt{n}} \leq z\right) - \Phi(z) \right| \leq \frac{4.08}{\sqrt{n}} (\|Y\|_6^6 + \|Z\|_6^6) (1 + \tilde{\sigma}^{-3}).$$

Recall:

here $V = (Y, Z, Y^2 - 1, Z^2 - 1, YZ)$;

so, $\|V\|_3^3 \asymp \|Y\|_6^6 + \|Z\|_6^6$.

Cf. Bhattacharya–Ghosh '78, Chibisov '80: Asymptotic expansion for the distr. of statistics admitting a stochastic expansion.

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7/23 Self-normalized sums: Bentkus and Götze, and Shao

Let Y, Y_1, \dots, Y_n be iid with $E Y = 0$ and $E Y^2 = 1$. Let

$$T := \frac{Y_1 + \dots + Y_n}{\sqrt{Y_1^2 + \dots + Y_n^2}} = \frac{\sqrt{n} \bar{Y}}{\sqrt{\bar{Y}^2}}.$$

Shao '05:

$$\begin{aligned} |\mathbb{P}(T \leq z) - \Phi(z)| &\leq 25 \frac{E |Y|^3}{\sqrt{n}} \mathbb{I} \left\{ |Y| \leq \frac{\sqrt{n}}{2} \right\} + 10.2 E Y^2 \mathbb{I} \left\{ |Y| > \frac{\sqrt{n}}{2} \right\} \\ &\leq 25 \frac{\|Y\|_3^3}{\sqrt{n}}; \end{aligned}$$

earlier, Bentkus and Götze '96: same without explicit constants.

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8/23 Central t and self-normalized sums

Chibisov '79–80; Slavova '85; Novak '00 and '05; Nagaev '02; and Pinelis '11 – specialized methods.

Pinelis '11:

$$|\mathbb{P}(T \leq z) - \Phi(z)| \leq \frac{1}{\sqrt{n}} \left(A_3 \|Y\|_3^3 + A_4 \|Y^2 - 1\|_2 + A_6 \frac{\|Y^2 - 1\|_3^3}{\|Y\|_3^3} \right)$$

for $(A_3, A_4, A_6) \in \{(1.53, 1.52, 1.34), (10.94, 9.40, 11.06 \times 10^{-6})\}$; the constants are slightly worse without the iid assumption.

Especially after truncation, this compares favorably with Shao's result.

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Pinelis and Molzon '11, based on a general result for statistics of form $f(\bar{V})$:

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10/23 Self-normalized sums, nonuniform BE bounds

$$|P(T \leq z) - \Phi(z)| \leq \frac{g(z)}{\sqrt{n}} (A_3 \|Y\|_3^3 + A_4 \|Y\|_4^8 + A_6 \|Y^2 - 1\|_3^3),$$

where $g(z) := \frac{1}{z^3} + \frac{w_g}{e^{z/2}}, \quad z \in (0, \omega\sqrt{n}]$,

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	$\omega = 0.1$			$\omega = 0.5$		
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$w_g = 0$	151	148	147	166	166	165
	170	85	29	229	115	45

Comment: $\max_{z>0} \frac{e^{-z/2}}{1/z^3} \approx 10.75$, attained at $z = 6$.

Cf.: best known nonuniform BE bound for sums of i.i.d. r.v.'s

(Michel '81 cum Shevtsova '11): $30.2211 \frac{\|Y\|_3^3}{(|z|^3+1)\sqrt{n}}$.

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11/23 Self-normalized sums vs. Student's statistic

For all $n > 1$

$$\left| \sup_{z \in \mathbb{R}} |\mathbb{P}(t \leq z) - \Phi(z)| - \sup_{z \in \mathbb{R}} |\mathbb{P}(T \leq z) - \Phi(z)| \right| < \frac{C}{n-1},$$

where t is the Student statistic, T is again the self-normalized sum,

$$C := \left(k - \frac{1}{2}\right) e^{-k} \sqrt{\frac{k}{\pi}} = 0.162\dots, \quad \text{and} \quad k := 1 + \frac{\sqrt{3}}{2}.$$

For all $n > 1$

$$\left| \sup_{z \in \mathbb{R}} |\mathbb{P}(t \leq z) - \Phi(z)| - \sup_{z \in \mathbb{R}} |\mathbb{P}(T \leq z) - \Phi(z)| \right| < \frac{C}{n-1},$$

where t is the Student statistic, T is again the self-normalized sum,

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12/23 Exact bounds on the closeness between t_p and $t_\infty = N(0, 1)$

Let $d_{\text{TV}}(p)$ and $d_{\text{Ko}}(p)$ denote, resp., the total-variation and Kolmogorov distances between t_p and $t_\infty = N(0, 1)$. Then

$$\frac{1}{2} d_{\text{TV}}(p) = d_{\text{Ko}}(p) < \frac{C}{p} \quad \forall p \in [4, \infty),$$

where

$$C := \frac{1}{4} \sqrt{\frac{7 + 5\sqrt{2}}{\pi e^{1+\sqrt{2}}}} = \lim_{p \rightarrow \infty} p d_{\text{Ko}}(p) = 0.158\dots,$$

so that C is the best possible factor.

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13/23 Tail monotonicity of t_p

Let G_p stand for the tail function of t_p . Then

$$0 < p < q \leq \infty \implies \frac{G_q(x)}{G_p(x)} \text{ is (strictly) decr. in } x \geq 0,$$

whence the stochastic majorization:

$$G_q(x) < G_p(x) \quad \forall x > 0.$$

Note: here the likelihood ratio is not monotone; so, the usual scheme $\text{MLR} \implies \text{MTR} \implies \text{SM}$ doesn't work.

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14/23 Exact upper bounds on the mean of the Winsorised-tilted distribution

For each $h > 0$ and each $w \in \mathbb{R}$, the maximum of the tilted mean

$$\frac{E X e^{h(X \wedge w)}}{E e^{h(X \wedge w)}}$$

given $E X$ and $E X^2$ is attained when X has a two-point distribution.

For $E X = 0$ and $w > 0$, this maximum is

$$< \frac{e^{hw} - 1}{w} E X^2,$$

and the factor $\frac{e^{hw}-1}{w}$ is the best possible.

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15/23 Exact lower bounds on the exponential moments of the Winsorised-tilted distribution

For each $h > 0$ and each $w > 0$, the minimum of

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given $E X \geq 0$ and $E X^2$ is attained when X has a two-point distribution.

For each $w > 0$, the minimum of these minima over all $h > 0$ is strictly positive (not so if $X \wedge w$ is replaced by $X I\{X \leq w\}$).

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16/23 An asymptotically Gaussian bound on the Rademacher tails

$$\begin{aligned}\mathbb{P}(a_1\varepsilon_1 + \cdots + a_n\varepsilon_n \geq x) &\leq \mathbb{P}(Z > x) + \frac{C\varphi(x)}{9 + x^2} \\ &< \mathbb{P}(Z > x) \left(1 + \frac{C}{x}\right) \quad \forall x > 0,\end{aligned}$$

where $\varepsilon_1, \dots, \varepsilon_n$ are independent Rademacher r.v.'s,

$$a_1^2 + \cdots + a_n^2 = 1,$$

$Z \sim N(0, 1)$ with density φ , and

$C := 5\sqrt{2\pi e} \mathbb{P}(|Z| < 1) = 14.10\dots$ is a best possible constant factor: the 1st inequality above turns into the equality when $x = n = 1$.

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17/23 Refined and generalized Bennett-Hoeffding bound

Let X_1, \dots, X_n be independent r.v.'s, with $S := X_1 + \dots + X_n$.

Take any σ, y , and β in $(0, \infty)$ s.t. $\varepsilon := \frac{\beta}{\sigma^2 y} \in (0, 1)$. Suppose that

$$\sum_i \mathbb{E} X_i^2 \leq \sigma^2, \quad \sum_i \mathbb{E} (X_i)_+^3 \leq \beta, \quad \mathbb{E} X_i \leq 0, \quad \text{and } X_i \leq y,$$

for all i . Then

$$\mathbb{E} f(S) \leq \mathbb{E} f(\eta_{\varepsilon, \sigma, y})$$

for all $f \in C^2$ s.t. f and f'' are nondecreasing and convex, where

$$\eta_{\varepsilon, \sigma, y} := \Gamma_{(1-\varepsilon)\sigma^2} + y \tilde{\Pi}_{\varepsilon\sigma^2/y^2},$$

$\Gamma_{a^2} \sim N(0, a^2)$, $\tilde{\Pi}_\theta := \Pi_\theta - \theta$, $\Pi_\theta \sim \text{Poisson}(\theta)$,

and Γ_{a^2} and Π_θ are independent.

Corollary: $\mathbb{P}(S \geq x) \leq \frac{2e^3}{9} \mathbb{P}^{\text{LC}}(\eta_{\varepsilon, \sigma, y} \geq x) \quad \forall x \in \mathbb{R}$,

where $\mathbb{R} \ni x \mapsto \mathbb{P}^{\text{LC}}(\eta \geq x)$ is the least log-concave majorant of $\mathbb{R} \ni x \mapsto \mathbb{P}(\eta \geq x)$.

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For p and s in $(0, \infty)$, $j \in \overline{-1, \lceil p - 1 \rceil}$, and any r.v. X with $E e^{sX} < \infty$ one has

$$E X_+^p = \frac{\Gamma(p+1)}{\pi} \int_0^\infty \Re e \frac{E e_j((s+it)X)}{(s+it)^{p+1}} dt,$$

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19/23 Improved & generalized von Bahr–Esseen inequality

$$\mathcal{F}_+^{1,2} := \{f \in C^1(\mathbb{R}) : f(0) = 0, f \text{ is even,} \\ f' \text{ is nondecreasing and concave on } [0, \infty)\};$$

e.g., $|\cdot|^p \in \mathcal{F}_+^{1,2} \quad \forall p \in (1, 2]$. Then

$$E f(S_n) \leq E f(X_1) + C_f \sum_{j=2}^n E f(X_j)$$

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Let X_1, \dots, X_n be independent r.v.'s with values in $\mathfrak{X}_1, \dots, \mathfrak{X}_n$, resp.; here, all spaces and functions are measurable. Suppose

$$Y := g(X_1, \dots, X_n)$$

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$$\begin{aligned} & | \mathbb{E} g(x_1, \dots, x_{i-1}, \tilde{x}_i, X_{i+1}, \dots, X_n) \\ & - \mathbb{E} g(x_1, \dots, x_{i-1}, x_i, X_{i+1}, \dots, X_n) | \leq \rho_i(\tilde{x}_i, x_i) \end{aligned}$$

holds for some functions $\rho_i: \mathfrak{X}_i \times \mathfrak{X}_i \rightarrow \mathbb{R}$ and all i , $x_j \in \mathfrak{X}_j$, and $\tilde{x}_i \in \mathfrak{X}_i$.

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Then $\forall f \in \mathcal{F}_+^{1,2} \setminus \{0\} \quad \forall x_i \in \mathfrak{X}_i$

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where

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Then for any $f \in \mathcal{F}^{2,3}$, any zero-mean r.v. X , and any $t \in \mathbb{R}$

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where $c := \frac{17+7\sqrt{7}}{27} = 1.315\dots$ is the best possible factor.

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